

1 Unifying individual and metapopulation scales with
2 stochastic population models: the effect of climate and
3 competition on tree range limits

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20 **Abstract**

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Despite recent calls to use demographic range models to scale the effect of individual dynamics in setting range limits, there is a growing body of evidence showing that tree species' performance is not correlated with their distribution. In this study, we ask whether the challenge in predicting species distribution from demographic rates stems from overlooking the inherent variability of forest systems and the underlying uncertainty of forest models. We use a stochastic Integral Projection Model to predict species-level intrinsic population growth for 31 eastern North American tree species. We introduce a novel metric for species-level performance we coined local suitable probability, which captures observed spatiotemporal stochasticity in climate and competition while accommodating model uncertainty. Our focus is on investigating how suitable probability changes across the cold-to-hot species range distribution over the mean annual temperature gradient. Our findings reveal a consistent, nearly linear decline in suitable probability from the cold to hot borders across the species. This change in suitable probability towards the orders is primarily driven by climate rather than competition. These results, supported by a novel approach accounting for uncertainty, enhance our understanding of the nuanced interplay between climate and competition across species ranges. We conclude by proposing a novel theory that uses the local suitable probability to establish a link between individual demographic rates and metapopulation dynamics.

Keywords: Integral Projection Models, Species distribution, Individual variability, Environmental stochasticity, Forest demography

1 Introduction

Climate warming poses a significant challenge for several species, particularly for trees that struggle to follow temperature warming and moving ranges (Sittaro et al. 2017). It is imperative to untangle the mechanisms governing their range limits to forecast how they will respond to climate change. The niche theory predicts that a species will be present in suitable environmental conditions that allow the

species to have a positive growth rate (Hutchinson 1957). From this theory, we can define the geographic distribution of a species as a manifestation of individual demographic rates, such as growth, survival, and recruitment (Holt 2009). By assuming these demographic rates change with the environment, we can predict a species' range limits based on its individuals' performance (Maguire Jr 1973, Holt 2009).

Biotic interaction is undoubtedly an essential driver of demographic rates and, thereby, a potential driver of range limits. A recent theoretical framework based on coexistence theory has been proposed to assess how biotic interactions can scale up to affect range limits (Godsoe et al. 2017). Formally, this framework evaluates the intrinsic population growth rate when the focal species is rare (Chesson 2000), both in scenarios where there is no competition (fundamental niche) and when competitive species reach equilibrium (realized niche). Numerous studies have explored the influence of climate and competition on the distribution of forest trees across their ranges. For instance, Ettinger and HilleRisLambers (2017) observed in field experiments that neighboring competition constrained individual performance within the range but facilitated better performance outside the range. Using a dynamic forest model, Scherrer et al. (2020) showed how slow demographic rates and negative competition reduce the uphill migration rate of 16 tree species. Despite this evidence, the application of this framework to predict the geographic distribution of species based on demographic rates often reveals weak correlations between the performance of tree species and their distribution (McGill 2012, Csergo et al. 2017, Bohnert and Diez 2020, Le Squin et al. 2021, Midolo et al. 2021, Guyennon et al. 2023, Thuiller et al. 2014).

One possible explanation for such discrepancy between demographic rates and species distribution is the common practice of assessing performance under average conditions and pointwise estimations, neglecting the associated uncertainty in these estimates. In ecological models, the uncertainty in estimation arises from three distinct sources. The first source involves measurement errors, a factor often neglected in ecological models (Damgaard 2020). The second is process uncertainty linked to model (mis)specification (Harwood and Stokes 2003), determined by all variation that is not captured by the model covariates. Finally, even with a well-defined model and precise data, models must also consider parameter uncertainty due to individual variability (Cressie et al. 2009, Shoemaker et al. 2020).

Beyond data and model uncertainty, variability in demographic rates and subsequently in the population growth rate (λ) arises from two primary sources (van Daalen and Caswell 2020). The first is attributed

74 to demographic and environmental stochasticity, where individuals exposed to identical conditions may
75 exhibit different responses simply by chance (Caswell 2009). The second source of variability arises
76 from heterogeneity encountered at various scales. These differences can manifest between individual
77 stages that motivated the development of structured population models (Lewis 1942, Leslie 1945),
78 and can promote high species diversity in forest trees (Clark 2010). Another source of heterogeneity
79 arises from large-scale differences in neighboring patches, often described by the metapopulation theory
80 (Levins 1969). This theory posits that the dynamics of occupied and empty patches in a landscape
81 are driven by colonization and extinction processes. Building on this theory, Talluto et al. (2017)
82 used patch variability to derive colonization and extinction rates of eastern North American trees,
83 revealing their distribution to be out of equilibrium with climate. Therefore, while this result advances
84 our understanding of the mechanisms governing large-scale tree distributions, there remains a need to
85 reconcile it with local demographic dynamics, given that colonization and extinction processes ultimately
86 manifest from demographic rates.

87 Theory predicts that the uncertainty arising from stochastic and heterogenous processes may lead to
88 divergent outcomes in λ . Demographic and environmental stochasticity may increase the uncertainty in
89 λ , consequently increasing the extinction risk, particularly for populations with low performance or low
90 density (Holt et al. 2005, Gravel et al. 2011). For instance, demographic stochasticity increased the
91 extinction risk of European forest trees at the hot edge of their distribution (Guyennon et al. 2023). On
92 the other hand, spatial heterogeneity has been described as a buffering process against the stochasticity
93 in demographic rates, thereby increasing population persistence (Milles et al. 2023). This is particularly
94 relevant in nonlinear models, where higher demographic and environmental stochasticity can increase
95 the difference in population growth rate compared to the average expectations (Koons et al. 2009).
96 Furthermore, demographic and environmental stochasticity influence abundance variation, indirectly
97 impacting λ through density-dependence (May et al. 1978, Terry et al. 2022). A comprehensive
98 understanding of the response of forest trees to climate change requires incorporating the multiple
99 sources of variability arising from spatio-temporal variation and parameter uncertainty.

100 Here, we use a stochastic Integral Projection Model (IPM) to predict species-level intrinsic population
101 growth (λ) for 31 eastern North American tree species. The IPM integrates the growth, survival, and
102 recruitment demographic rates, which vary in response to climate and competition. By fitting each

103 demographic rate using non-linear hierarchical Bayesian models, we capture parameter uncertainty
104 at both the individual and local population scales. Additionally, our model naturally accommodates
105 observed spatio-temporal stochasticity in climate and competition. Then, rather than ignoring these
106 sources of variability, we embrace them into λ by defining species performance through a probabilistic
107 framework. Specifically, we introduce a novel metric called **local suitable probability**, derived from
108 the average population growth rate and its associated variability. This metric determines the probability
109 of a positive population growth rate for a species under specific climate and competition conditions.

110 In our analysis, we first used the IPM to predict species-specific λ at the plot level under two conditions:
111 without (fundamental niche) and with (realized niche) heterospecific competition. We replicated this
112 calculation 100 times across all observed plots from the same species to assess the variability of λ
113 arising from both spatio-temporal stochasticity in the climate and competition and model uncertainty.
114 As this variable λ changes across space, we used these observations to model how the species' local
115 suitable probability changes across the mean annual temperature. Specifically, we ask how climate and
116 competition affect each species' local suitable probability. Then, we investigated how a species' local
117 suitable probability changes from the center of its distribution toward the cold and hot borders. Finally,
118 we disentangle the relative impacts of climate and competition in changing suitable probability from the
119 center to the borders. We conclude by discussing a novel theory that uses the local suitable probability
120 to establish a link between individual demographic rates and metapopulation dynamics.

121 2 Methods

122 2.1 Population model, demographic components, and uncertainty 123 structure

124 We use an Integral Projection Model (IPM) to predict the intrinsic population growth rate (λ) as
125 a function of climate and competition. An IPM is a powerful modeling approach that allows a full
126 representation of all sources of variability in demography. The IPM serves as a mathematical formulation
127 describing the dynamics of a continuous trait distribution (z) within a population over discrete time
128 steps:

$$n(z', t + 1) = \int_L^U K(z', z, X, \theta) n(z, t) dz \quad (1)$$

129 In our case, the trait z is defined as the tree's diameter at breast height (dbh), constrained between the
 130 limits L and U . The continuous distribution $n(\cdot)$ of dbh z of a population at time t transitions to the
 131 next time step using a projection kernel (K). The kernel K , with parameters θ and covariates X that
 132 are time dependent, comprises three demographic submodels:

$$K(z', z, \theta) = [Growth(z', z, X, \theta) \times Survival(z, X, \theta)] + Recruitment(z, X, \theta) \quad (2)$$

133 The growth model assesses the probability of an individual of size z at time t transitioning to size z' at
 134 time $t + 1$. The survival model determines the probability of an individual with size z at time t surviving
 135 to the next time step. Lastly, the recruitment model determines the number of new individuals entering
 136 the population at each time step as a function of total density z . The kernel K has the same function
 137 of the population growth rate r in a population model, where multiplying the population distribution
 138 $n(z, t)$ with K gives the population distribution at the next time step $n(z', t + 1)$. Its advantage in
 139 propagating uncertainty is that, instead of having a matrix with fixed parameters determining the
 140 transition rate of population individuals over time, it uses a probability distribution with uncertainty
 141 derived from the demographic models to project individuals over time.

142 With the defined K , we can estimate the intrinsic population growth rate for a determined set of
 143 conditions from the covariates X and sampled parameters from the posterior distribution θ . Specifically,
 144 we discretize the continuous kernel K using the mid-point rule (Ellner et al. 2016) and estimate the
 145 intrinsic population growth rate using the dominant eigenvalue of the discretized K . This approach is a
 146 local approximation of the population growth rate at the initial time steps.

147 A detailed description of the data and model development is available in Chapter 2. In summary, we
 148 evaluated non-linear statistical models to formulate the growth, survival, and recruitment components
 149 of the IPM, along with their uncertainty. Each demographic sub-model varies as a function of the
 150 mean annual temperature, mean annual precipitation, and stand basal area of larger individuals. Each

151 model's parameters (θ) are species-specific, as each model is fitted separately for each species. Both
152 climate variables influence each demographic model through an unimodal link function, where each
153 model exhibits an optimal climate and niche breadth for temperature and precipitation. Additionally,
154 density dependence is integrated based on the plot's total basal area of larger individuals. Stand density
155 affects growth and survival through a linear model, in which two parameters determine the strength of
156 interaction from conspecific and heterospecific (all species combined) competition. For the recruitment
157 model, the annual ingrowth rate is modulated by conspecific stand basal area, using an unimodal
158 function to account for both the positive effect of seed source and the negative effect of conspecific
159 competition. Furthermore, the annual survival rate of potential ingrowth individuals decreases linearly
160 with the stand density of heterospecific individuals. Finally, the intercept of each growth, survival, and
161 recruitment model incorporates plot-level random effects to control for the variance shared within the
162 plot-year observations.

163 We use two open inventory datasets from eastern North America: the Forest Inventory and Analysis
164 (FIA) dataset in the United States (O'Connell et al. 2007) and the permanent plots of forest inventory
165 program for Québec (Ministère des Ressources Naturelles 2016). These inventories, with multiple
166 individual measurements over time and space, allowed us to use the transition information between
167 measurement years for predicting growth, survival, and recruitment rates. We selected the 31 most
168 abundant species, comprising 9 conifer species and 22 hardwood species, well-dispersed across shade
169 tolerance and successional status (Supplementary Material 1). These species are well distributed across
170 the eastern North American gradient and the sampling area covers cold and hot range limits for most
171 species.

172 2.2 Extracting local suitability probability

173 We estimate λ at the local population scale, specifically at the plot level in our study. Within a given
174 geographic location, such as a specific latitude where several plots are located, the variance of λ among
175 those plots arises from spatio-temporal variations in both climate and competition covariates. For
176 instance, climate stochasticity introduces noise in annual temperature and precipitation, leading to
177 environmental variation. Similarly, even with identical climate conditions, two locations can exhibit
178 different community abundance and composition, resulting in variability in the strength of competition.

179 Beyond these spatio-temporal environmentally-induced variations, λ can still vary due to the other
180 sources of uncertainty discussed above.

181 We track demographic model uncertainty at two complementary scales: individual and plot levels. At
182 the individual level, without plot random effects, two plots with the same climate and competition
183 conditions may have different λ values due to the uncertainty in the demographic growth, survival, and
184 recruitment sub-models. Similarly, with the same environmental conditions and averaged parameter
185 values (eliminating demographic uncertainty at the individual level), two plots can still yield different
186 λ values due to the spatial uncertainty of each demographic model due to the plot random effects.
187 Therefore, variability in the population growth rate can arise from spatio-temporal variations in both
188 the environment and the parameters.

189 Given these different sources of variability in λ , we define the suitable probability as the area under the
190 distribution for $\lambda \geq 1$. To estimate this, we first determine the cumulative distribution function, $F(x)$,
191 from the generic probability density function, $\lambda = f(t)$, as follows:

$$F_{\lambda}(x) = P(\lambda \leq x) = \int_{-\infty}^x f(t)dt \quad (3)$$

192 This function represents the cumulative distribution from $-\infty$ to x . Subsequently, we define the suitable
193 probability (Λ) as the complement of the cumulative distribution function for $x = 1$:

$$\Lambda = 1 - F_{\lambda}(1) \quad (4)$$

194 2.3 Modeling suitable probability

195 We can evaluate the suitable probability of a species at various scales, ranging from a single local
196 plot up to several plots in a region. At the plot level, sources of variability in λ stem from parameter
197 uncertainty, individual heterogeneity, and temporal variability in climate and competition. When
198 considering multiple plots simultaneously, we can additionally account for spatial variability in climate
199 and competition, along with spatial uncertainty in plot-level parameters.

200 Apart from parameter uncertainty at the individual level, all other sources of variability exhibit spatial
 201 dependence. This implies that environmental variability (from climate, competition, or both) and
 202 parameter uncertainty at the plot level can vary based on their spatial location. For instance, plots at
 203 the border of the species distribution may experience more temperature variability than those at the
 204 center. Additionally, plot-level parameter uncertainty can be spatially clustered, capturing potential
 205 features of demographic variability beyond the climatic and competition covariates, such as historical
 206 factors or local edaphic conditions.

207 Given that variability can be spatially dependent, we can model how suitable probability changes across
 208 the species' range distribution, considering both fixed climate and competition effects and the underlying
 209 spatio-temporal variability. We are particularly interested in how suitable probability changes from the
 210 center toward the cold and hot ranges. For that, we categorized all species' plot-year observations based
 211 on the gradient of mean annual temperature (MAT), divided into cold and hot ranges using the MAT
 212 centroid among all plots for the species ($\frac{\max(MAT) + \min(MAT)}{2}$). For instance, if a species is observed
 213 within the 4-10°C gradient of MAT, the plots with MAT below 7°C are classified as cold, while the
 214 others are classified as hot. We chose to use MAT instead of latitude because we are interested in the
 215 species' climatic niche, although the two variables are highly correlated.

216 We assessed suitable probability separately for the cold and hot ranges, employing a linear model to
 217 determine the relationship between λ and MAT. The spatio-temporal variability of λ arising from
 218 environmental stochasticity and parameter uncertainty influences the variance of the linear model. As
 219 this variance may change depending on the range position, we introduce a submodel for the variance of
 220 the linear model to be dependent on MAT. To accommodate potential asymmetry in this variance, we
 221 use a Skew Normal Distribution (SN) incorporating an additional parameter (α) that can introduce
 222 right or left-skewed tails to the variance:

$$\begin{aligned}
 \log(\lambda) &\sim SN(\xi, \omega, \alpha) \\
 \xi &= \beta_{1,\xi} \times MAT + \beta_{0,\xi} \\
 \omega &= e^{\beta_{1,\omega} \times MAT + \beta_{0,\omega}}
 \end{aligned} \tag{5}$$

223 Here, ξ is the location parameter or the λ average, and ω is the scale representing the variance around
224 the mean.

225 2.3.1 Simulations

226 We computed λ for each species based on the plot-year observations in the dataset, considering both
227 environmentally induced variability and parameter uncertainty. For every observed species-plot-year
228 combination, we incorporated temporal stochasticity in climate conditions by using the mean and
229 standard deviation of mean annual temperature and precipitation calculated from the years between
230 measurements. For instance, in the case of a plot observed twice, we calculated λ for the second
231 observation with climate conditions drawn randomly from a normal distribution with mean and standard
232 deviation extracted from plot specific climate observations for each year within the time interval.
233 Similarly, temporal stochasticity in competition arises from variation in abundance and composition
234 between measured years. By iteratively performing this calculation, drawing parameter values randomly
235 from the posterior distribution, we introduced demographic uncertainty at the individual level. For each
236 species-plot-year measurement, we replicated the calculation of λ 100 times. By applying this approach
237 across all plots, we naturally incorporate spatial variation in climate and competition conditions and
238 spatial uncertainty in plot-level parameters.

239 For each species-plot-year-replication combination, we calculated λ under two simulated conditions.
240 The first scenario excludes competition in order to evaluate the fundamental niche, with heterospecific
241 competition set at zero and conspecific total population size (N) set at 0.1. This simulation is used to
242 assess the fundamental niche. The second scenario is used to evaluate the invasion growth rate with
243 residents (the realized niche), with an evaluation of the population growth rate when the focal species
244 is rare ($N = 0.1$) and heterospecific competition is set to the observed abundance of the competitive
245 species. This condition simulated the population growth rate under the realized niche.

246 We then fitted a linear model of λ for each species-plot-year-replication as a function of the mean annual
247 temperature gradient. Species-specific linear models were evaluated for the hot and cold ranges using
248 the Hamiltonian Monte Carlo (HMC) algorithm via the Stan software (version 2.30.1 Team and Others
249 2022) and the `cmdstandr` R package (version 0.5.3 Gabry et al. 2023). We conducted 1000 iterations
250 for the warm-up and 1000 iterations for the sampling phase for each of the four chains, resulting in 4000

posterior samples (excluding the warm-up). We used a sample of 5000 plots for each species to fit the model. This sample was necessary only for 6 out of the 31 species.

We leveraged the posterior distribution to estimate the suitable probability of a species for any value of MAT under fundamental or realized niches for the cold and hot ranges. Specifically, we estimated suitable probability under four different MAT conditions encountered by the species: at the border and the center of each cold and hot range. We defined the border of the cold range as the minimum observed MAT for the focal species in the dataset, while the hot range was defined as the maximum observed MAT. The center location is defined as the centroid of MAT for the focal species. Although the center location has the same MAT for the cold and hot ranges, both are retained because the model is fitted separately for the cold and hot ranges. Finally, we estimated suitable probability for each location under no competition (fundamental niche) and heterospecific competition (realized niche) conditions, using the empirical cumulative distribution function over 1000 predictive draws.

The code for the computation of each plot-year λ is available at https://github.com/willvieira/forest-IPM/tree/master/simulations/lambda_plot, and the code to model the linear model is at https://github.com/willvieira/forest-IPM/tree/master/simulations/model_lambdaPlot/.

3 Results

3.0.1 Model fit

We first analyzed how the local population growth rate (λ) and its variability change across the cold and hot ranges (Equation 5). An example is provided at Figure 1 with the observed distribution of λ and the fit of the underlying model on the mean annual temperature gradient for balsam fir, *Abies balsamea*. Each point represents a plot-year-replication encompassing the complete spatio-temporal sources of variability arising from the stochastic environment and parameter uncertainty. The black line represents the fitted model of how λ changes with MAT, and the envelopes depict the 90th quantiles of model distribution. From this uncertainty, we can deduce the suitable probability. This example shows that the mean and variance of λ decrease towards the cold border, while it does not vary much towards the hot border. By comparing the model under heterospecific competition with that without

277 competition for the cold range, we observed that while their average is similar, the uncertainty of the
 278 model under heterospecific competition shifted downwards (Figure 1, bottom left).

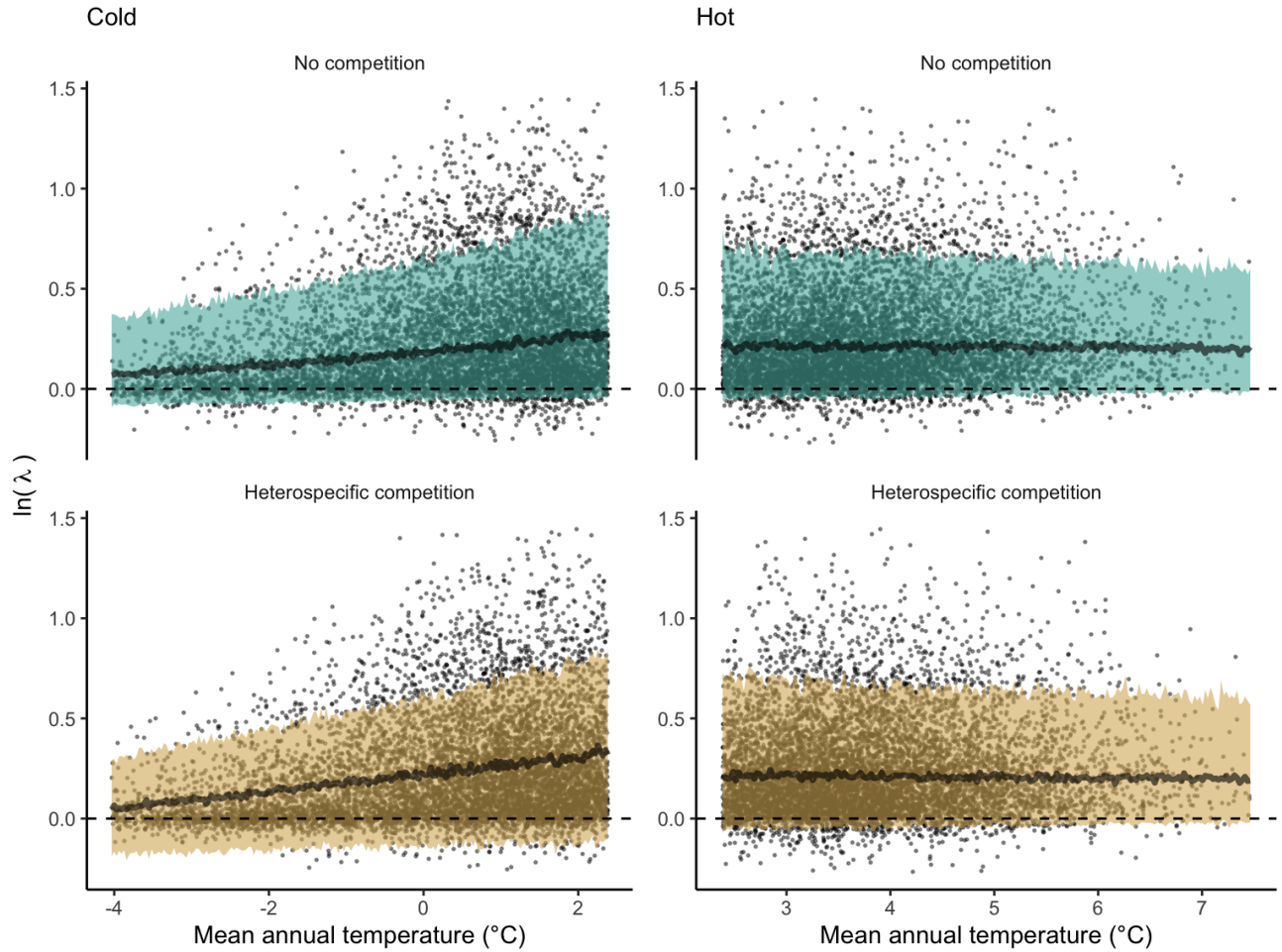


Figure 1: Distribution of stochastic population growth rate (λ) for *Abies balsamea* over the mean annual temperature gradient for different conditions. Species' population are split into cold (left panels) and hot (right panels) ranges under no competition (fundamental niche) and heterospecific competition (realized niche). The dots represent λ over the plot-year-replication combinations. The model's average line and 90% prediction intervals are estimated using 500 draws from the posterior distribution.

279 We then investigated the local suitability probability using the empirical cumulative distribution approach
 280 (Equation 4) from the linear model predictions. The Figure 2 shows the suitable probability expected
 281 over the mean annual temperature of the same species. We observed that the local suitability probability
 282 was reduced towards the cold border, with a stronger reduction under heterospecific competition (yellow
 283 curve). We can also observe that the decrease in suitable probability towards the border is nonlinear,
 284 becoming more substantial for heterospecific competition than for the no-competition condition.

285 The model fit and the estimation of suitable probability across the temperature gradient for all species
 286 are presented in Supplementary Material 2. We observed for most species a decrease of the climate
 287 effect at one border while the other remained unchanged. Additionally, a few species displayed a clear
 288 linear pattern of decreasing suitable probability from the cold to the hot border, with only one species
 289 (*Betula papyrifera*) having a decrease at both borders. Conversely, under the competition effect, most
 290 species exhibited a decrease in suitable probability at the hot border and an increase at the cold border,
 291 indicating a linear rise in the impact of competition from the cold to the hot border of the distribution.

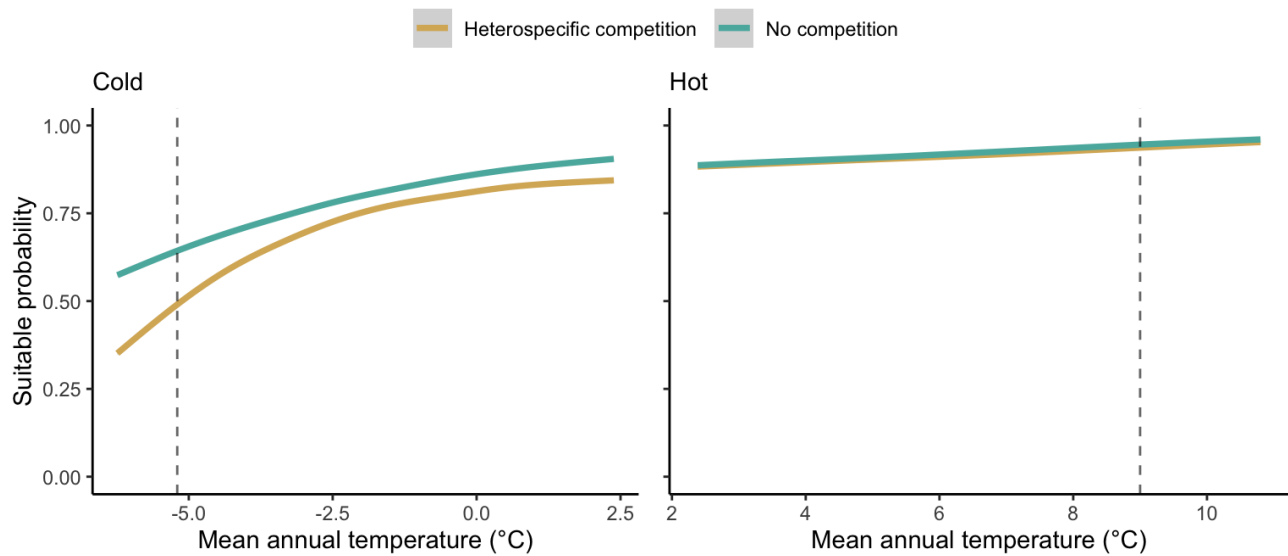


Figure 2: Suitable probability of *Abies balsamea* over the mean annual temperature gradient for cold and hot ranges under no competition (green) and heterospecific (yellow). The vertical dotted line represents the range limits of the MAT observed in the dataset.

292 3.0.2 Effect of climate and competition on suitable probability for the center and 293 border distributions

294 We investigated the effect of climate and competition on the suitable probability at the border and
 295 center of the temperature range distribution for all species. Because the border and center positions are
 296 relative to each species, we could not represent the continuous trend in suitable probability across the
 297 MAT for all 31 species together. Instead, we extracted the local suitable probability with and without
 298 heterospecific competition for four locations across the MAT gradient (Figure 3). Overall, suitable
 299 probability was high among the species, with an average of 0.78. Among the four locations, species
 300 presented a lower suitable probability at the border of the hot range, with an average of 0.67. Across the

301 temperature range, there is a monotonic decrease in suitable probability from the cold border toward
 302 the hot border.

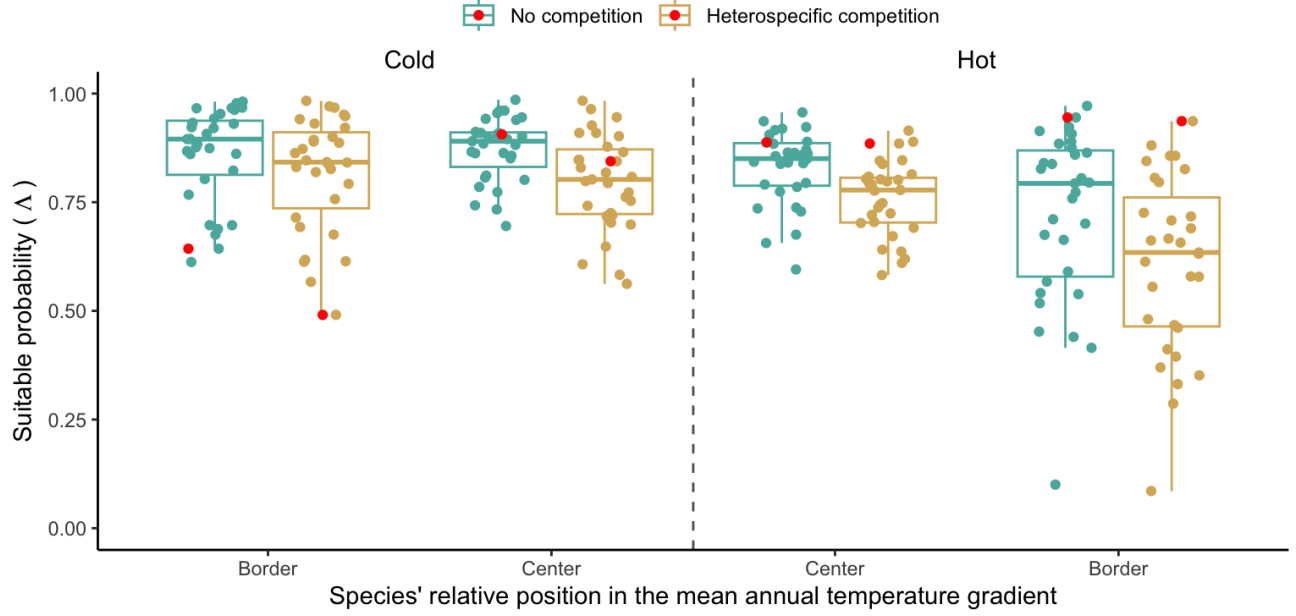


Figure 3: Estimated suitable probability for the 31 forest species across the center and border of the cold and hot ranges. The x-axis represents the mean annual temperature gradient similar to Figure 2, but is discretized at the border and center limits relative to each species. We highlighted the balsam fir species in red. Note that we omitted the parameter uncertainty of each species in this figure to avoid overlap and increase clarity.

303 We further disentangle the influence of competition from that of climate by calculating the difference
 304 between suitable probability under heterospecific competition and without competition. A negative
 305 difference signifies competition reduces suitable probability, while positive differences indicate an increase.
 306 Across the four climate locations, heterospecific competition consistently reduced suitable probability
 307 for most species, with the magnitude of reduction intensifying from the cold to the hot border (Figure
 308 S1). This suggests that the decline in suitable probability observed from the cold to the hot border
 309 (Figure 3) results from the combined effect of climate and competition.

310 3.0.3 Suitable probability change from center to border

311 We investigated the relative effect of climate and competition on changing suitable probability from the
 312 center to the border of the species distribution (Figure 4). A positive relative difference indicates an
 313 increase in suitable probability from the center towards the border, while a negative difference indicates

314 a decrease. Most species exhibited a decrease in suitable probability at the hot border relative to the
 315 center. Alternatively, most species showed a reduction in the effect of competition toward the cold
 316 border. However, the climate effect in the cold range was more variable, with some species experiencing
 317 an increase and others a decrease in suitable probability (Figure S2). Overall, the relative difference in
 318 suitable probability from the center toward the cold and hot borders was more influenced by climate
 319 rather than competition.

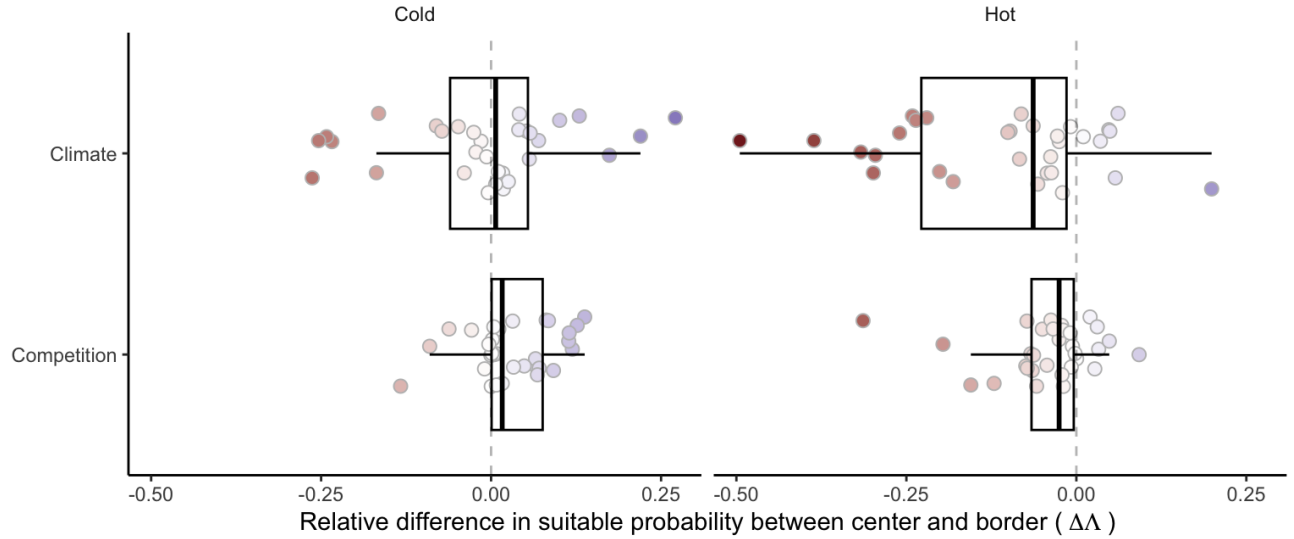


Figure 4: Difference in suitable probability for climate and competition effects over the cold and hot ranges. Negative values denote a decrease in species suitable probability from the center towards the distribution border, while positive values indicate an increase. Specifically, a negative value for climate at the hot (or cold) range signifies a reduction in suitable probability as temperature rises (or falls) towards the border. Boxplots determine the 25-75 quantile distribution among the species.

320 4 Discussion

321 Understanding the mechanisms shaping species distribution is imperative to face ongoing global changes.
 322 We acknowledged and integrated various sources of variability in the population growth rate of forest
 323 trees, contributing to an improved understanding of forest dynamics in an uncertain world. Introducing
 324 a novel metric, we quantified the relative impacts of climate and competition on the change in suitable
 325 probability across species distributions. Our findings revealed a nearly linear reduction in suitable
 326 probability from the cold to hot borders. Notably, the predominant influence on the relative difference in
 327 suitable probability from the center toward the border was attributed to climate rather than competition.

328 These results, supported by a novel approach accounting for uncertainty, enhance our understanding of
329 the nuanced interplay between climate and competition across species ranges.

330 The suitable probability was high across all species and range locations, with only around 5% of all
331 species-location combinations having a suitable probability below the 0.5 threshold. This is primarily
332 attributed to most species exhibiting a high positive population growth rate across their current
333 range distribution. Additionally, the spatio-temporal variability in the environment and the parameter
334 uncertainty in the plot may contribute to the elevated average population growth rate due to nonlinear
335 averaging. This aligns with theoretical (Schreiber and Lloyd-Smith 2009) and empirical (Crone 2016)
336 studies suggesting that spatial heterogeneity should increase the population growth rate.

337 Competition significantly reduces local suitability across all range locations, with a stronger and more
338 consistent effect at the cold border, contributing to the ongoing debate surrounding its significance
339 in setting range limits. Despite several studies emphasizing the effect of competition compared to
340 climate on the demographic rates of forest trees (Zhang et al. 2015, Käber et al. 2021, Le Squin et al.
341 2021), debates persist regarding whether this effect at the local scale translates to the biogeographic
342 distribution of species (Soberón 2007, Copenhaver-Parry et al. 2017). Our findings support the Godsoe
343 et al. (2017) hypothesis and a growing body of evidence (Scherrer et al. 2020, Shi et al. 2020, Paquette
344 and Hargreaves 2021, Lyu and Alexander 2022) showing that the effect of competition on the intrinsic
345 population growth rate can indeed contribute to range limits.

346 The decline in suitable probability from the cold to the hot border suggests a predominantly linear, rather
347 than unimodal, relationship with temperature for most species (Figure S3). This result is consistent
348 with reduced population growth rates in North American (Le Squin et al. 2021, Schultz et al. 2022) and
349 European (Guyennon et al. 2023) forest trees, except for the contrasting pattern observation by Purves
350 (2009). The higher suitable probability in the cold range compared to the hot range could be attributed
351 to multiple factors. First, species may still follow their climate niche post the last glaciation, explaining
352 why the current cold range limit does not align with the expected niche distribution (Svenning and
353 Skov 2007), potentially leading to a colonization debt (Talluto et al. 2017). Notably, four of the six
354 species exhibiting a significant decrease in suitable probability from the center toward the cold range
355 were already at the extreme cold observed in the dataset (Figure S4).

Our model may however overlook crucial drivers of species performance, despite capturing a substantial amount of variation from parameter uncertainty at the plot level. Factors such as the impact of extreme temperature and precipitation on phenology can influence tree range limits (Morin et al. 2007). Beyond covariates and plot-level uncertainty, incorporating temporal uncertainty at the plot level, accounting for spatio-temporal covariance, could likely capture additional sources of variation in demographic rates. While our approach considers temporal stochasticity in climate and competition, which affect species range size (Holt et al. 2022), there remains temporal variation in demographic rates beyond these covariates. This variability, possibly captured with random effects at the plot level, can influence range limits based on the degree of temporal autocorrelation and its relationship with the range (Benning et al. 2022). For instance, an empirical study on perennial herbaceous species demonstrated that temporal environmental stochasticity reduced the population growth rate relative to the average (Crone 2016). In our study, this temporal variability is particularly relevant for survival (due to disturbance) and recruitment (due to phenology) rates because, in addition to having high temporal variability (Clark et al. 1999, de Souza Leite et al. 2023), they represent the most significant drivers of population growth rate (Chapter 2).

The effect of competition, similar to climate, increased from the center towards the border of the hot range, contrary to Kunstler et al. (2021), who found no difference in the competition effect between the center and border of the species. Additionally, our results deviate from the Species Interactions-Abiotic Stress Hypothesis, predicting a stronger competition effect in less stressful climate conditions (Louthan et al. 2015). When considering the relative position of the species across the temperature gradient, only the effect of climate at the cold range changed with temperature. This indicates that most species have a similar or higher suitable probability at the border of the cold range compared to their center distribution. We further tested whether the species' range size affects the relative difference in suitable probability; while the absolute values change, the pattern among the species remains unchanged.

The climate gradient of temperature had a more significant effect than competition in changing the suitable probability of forest trees. This means that mean annual temperature, along with all latent variables, better explains how suitable probability changes across the temperature range. The choice of using only mean annual temperature as an explanatory variable for the variance of λ can be improved. For instance, the model could be built accounting for mean annual temperature and precipitation to

385 predict the complete two-dimensional distribution of the species' climate niche. Plot random effects
 386 could be further used to account for the nestedness of the data design, allowing the proper separation of
 387 the total variance of the metamodel into variance arising from individual- and plot-level demographic
 388 uncertainty. While we have assumed climate variability as independent and identically distributed
 389 random variables, this assumption can be relaxed to include temporal autocorrelation. Autocorrelated
 390 environmental fluctuation can significantly change a species' range limits due to nonlinear averaging
 391 (Benning et al. 2022, Holt et al. 2022). Lastly, although coexistence theory assumes the abundance of
 392 competitors to be at equilibrium (Chesson 2000), testing this assumption remains practically impossible.

393 Despite the many ways of improving our study, there is a growing body of evidence indicating a mismatch
 394 between performance and occurrence (McGill 2012, Csergo et al. 2017, Bohner and Diez 2020, Le Squin
 395 et al. 2021, Midolo et al. 2021, Guyennon et al. 2023, Thuiller et al. 2014). Our approach can better
 396 capture the nuanced effect of climate and competition along with the spatio-temporal variation in λ ,
 397 yet it was not enough to fully predict tree range limits. Since species distribution is influenced by
 398 processes at multiple scales (McGill 2010, Heffernan et al. 2014), it is challenging to rely on a single
 399 individual-level performance metric to predict it all (Evans et al. 2016). For instance, dispersion plays a
 400 crucial role in changing species distribution at larger spatial scales, either reducing its extent due to
 401 limited dispersal or increasing it through source-sink dynamics (Pulliam 2000). We propose that our
 402 novel metric, local suitable probability, can be a key unifying factor linking local and landscape scales.

403 Forest trees exhibit variation in their frequency of occurrence across distribution gradients, yet their
 404 relative abundance remains consistent when present (Canham and Thomas 2010). Such observation
 405 implies that assessing forest distribution should focus on colonization and extinction patch dynamics
 406 rather than local performance (Canham and Murphy 2017). However, instead of restricting models
 407 to either local or large scales, we propose using the local suitable probability to reconcile the local
 408 demographic dynamics with the metapopulation theory. Colonization and extinction processes, as
 409 described by metapopulation theory (Levins 1969), are well-suited for describing the mosaic of forest
 410 successional stages at the landscape scale resulting from natural disturbances and succession. However,
 411 an implicit assumption is that unoccupied patches are necessarily available for colonization. We relax this
 412 assumption and quantify patch availability using the local suitable probability metric (Λ). Considering an
 413 ensemble of patches (p) where individuals can arrive and establish in empty patches through colonization

(α), and occupied patches can become empty through extinction (ε), the integrated metapopulation model becomes:

$$\frac{dp}{dt} = \alpha p(\Lambda - p) - \varepsilon p \quad (6)$$

With this formulation, rather than having $1 - p$ available patches for colonization, we have $\Lambda - p$. Therefore, when λ and its variability are high, the local suitable probability equals 1, indicating that all non-occupied patches are available. Conversely, as the local suitable probability decreases, the proportion of non-occupied patches available for colonization is reduced. This integrative approach allows one to account for both the local (e.g. competition and climate) and landscape (e.g. fire disturbances and dispersal) drivers of forest dynamics when assessing tree distribution.

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